

Confirmatory factor analysis and Raykov's rho

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Contents

1	Introduction	1
2	Preliminaries	2
2.1	Load libraries	2
2.2	Load data set	2
3	Confirmatory factor analysis	3
3.1	Preliminary steps	3
3.2	Step 1	5
3.3	Step 2	5
3.4	Step 3	10
4	Construct reliability	13
5	Path diagram	14
6	Results presentation	15
	References	17

1 Introduction

In this hands-on, we are going to further validate our model based on the EFA findings. The same data set, "Attitude_Statistics v3.sav" will be used.

The evidence of internal structure will be provided by

1. Confirmatory factor analysis
 - Model fit
 - Factor loadings
 - Factor correlations (no multicollinearity)
2. Construct reliability
 - Raykov's rho

2 Preliminaries

2.1 Load libraries

In addition to `psych` (Revelle, 2019), we are going to use `lavaan` version 0.6.3 (Rosseel, 2018), `semTools` version 0.5.1 (Jorgensen, Pornprasertmanit, Schoemann, & Rosseel, 2018) and `semPlot` version 1.1.1 (Epskamp, 2019) in our analysis. `ls"")` [`@R-semTools`] `andsemPlot`' version 1.1.1 (Epskamp, 2019) in our analysis. These packages must be installed from downloaded packages from CRAN at <https://cran.r-project.org/>.

Again, make sure you already installed all of them before loading the packages.

```
library(foreign)
library(psych)
library(lavaan) # for CFA
library(semTools) # for additional functions in SEM
library(semPlot) # for path diagram
```

2.2 Load data set

We include only good items from **PA1** and **PA2** in `data.cfa` data frame.

```
data = read.spss("Attitude_Statistics v3.sav", F, T) # Shortform
# Include selected items from PA1 & PA2 in "data.cfa"
data.cfa = data[c("Q4", "Q5", "Q6", "Q7", "Q8", "Q9", "Q10", "Q11")]
str(data.cfa)

## 'data.frame': 150 obs. of 8 variables:
## $ Q4 : num 3 3 1 4 2 3 3 2 4 4 ...
## ..- attr(*, "value.labels")= Named num 5 4 3 2 1
## ... - attr(*, "names")= chr "STRONGLY AGREE" "AGREE" "NEUTRAL" "DISAGREE" ...
## $ Q5 : num 4 4 1 3 5 4 4 3 3 4 ...
## ..- attr(*, "value.labels")= Named num 5 4 3 2 1
## ... - attr(*, "names")= chr "STRONGLY AGREE" "AGREE" "NEUTRAL" "DISAGREE" ...
## $ Q6 : num 4 4 1 2 1 4 3 3 4 5 ...
## ..- attr(*, "value.labels")= Named num 5 4 3 2 1
## ... - attr(*, "names")= chr "STRONGLY AGREE" "AGREE" "NEUTRAL" "DISAGREE" ...
## $ Q7 : num 3 4 1 2 4 4 4 2 3 4 ...
## ..- attr(*, "value.labels")= Named num 5 4 3 2 1
## ... - attr(*, "names")= chr "STRONGLY AGREE" "AGREE" "NEUTRAL" "DISAGREE" ...
## $ Q8 : num 3 3 4 2 5 3 3 3 3 4 ...
## ..- attr(*, "value.labels")= Named num 5 4 3 2 1
## ... - attr(*, "names")= chr "STRONGLY AGREE" "AGREE" "NEUTRAL" "DISAGREE" ...
## $ Q9 : num 3 3 4 2 5 4 3 4 5 4 ...
## ..- attr(*, "value.labels")= Named num 5 4 3 2 1
## ... - attr(*, "names")= chr "STRONGLY AGREE" "AGREE" "NEUTRAL" "DISAGREE" ...
## $ Q10: num 3 3 5 2 3 4 3 4 4 4 ...
## ..- attr(*, "value.labels")= Named num 5 4 3 2 1
## ... - attr(*, "names")= chr "STRONGLY AGREE" "AGREE" "NEUTRAL" "DISAGREE" ...
## $ Q11: num 4 4 1 3 4 4 3 3 4 4 ...
## ..- attr(*, "value.labels")= Named num 5 4 3 2 1
## ... - attr(*, "names")= chr "STRONGLY AGREE" "AGREE" "NEUTRAL" "DISAGREE" ...

dim(data.cfa)

## [1] 150 8
```

```

names(data.cfa)

## [1] "Q4"  "Q5"  "Q6"  "Q7"  "Q8"  "Q9"  "Q10" "Q11"

head(data.cfa)

##   Q4 Q5 Q6 Q7 Q8 Q9 Q10 Q11
## 1  3  4  4  3  3  3   3   4
## 2  3  4  4  4  3  3   3   4
## 3  1  1  1  1  4  4   5   1
## 4  4  3  2  2  2  2   2   3
## 5  2  5  1  4  5  5   3   4
## 6  3  4  4  4  3  4   4   4

```

3 Confirmatory factor analysis

3.1 Preliminary steps

Descriptive statistics

Check minimum/maximum values per item, and screen for any missing values,

```

describe(data.cfa)

##    vars   n  mean    sd median trimmed  mad min max range skew kurtosis    se
## Q4     1 150 2.81 1.17      3  2.77 1.48   1   5     4  0.19 -0.81 0.10
## Q5     2 150 3.31 1.01      3  3.32 1.48   1   5     4 -0.22 -0.48 0.08
## Q6     3 150 3.05 1.09      3  3.05 1.48   1   5     4 -0.04 -0.71 0.09
## Q7     4 150 2.92 1.19      3  2.92 1.48   1   5     4 -0.04 -1.06 0.10
## Q8     5 150 3.33 1.00      3  3.34 1.48   1   5     4 -0.08 -0.12 0.08
## Q9     6 150 3.44 1.05      3  3.48 1.48   1   5     4 -0.21 -0.32 0.09
## Q10    7 150 3.31 1.10      3  3.36 1.48   1   5     4 -0.22 -0.39 0.09
## Q11    8 150 3.35 0.94      3  3.37 1.48   1   5     4 -0.31 -0.33 0.08

```

Note that all $n = 150$, no missing values. `min-max` cover the whole range of response options.

% of response to options per item,

```

response.frequencies(data.cfa)

##       1     2     3     4     5 miss
## Q4 0.140 0.280 0.30 0.19 0.093   0
## Q5 0.040 0.167 0.35 0.33 0.113   0
## Q6 0.080 0.233 0.33 0.26 0.093   0
## Q7 0.133 0.267 0.23 0.29 0.080   0
## Q8 0.047 0.100 0.48 0.23 0.147   0
## Q9 0.047 0.093 0.42 0.25 0.187   0
## Q10 0.073 0.107 0.42 0.23 0.167   0
## Q11 0.027 0.153 0.35 0.39 0.087   0

```

All response options are used, and there are no missing values.

Multivariate normality

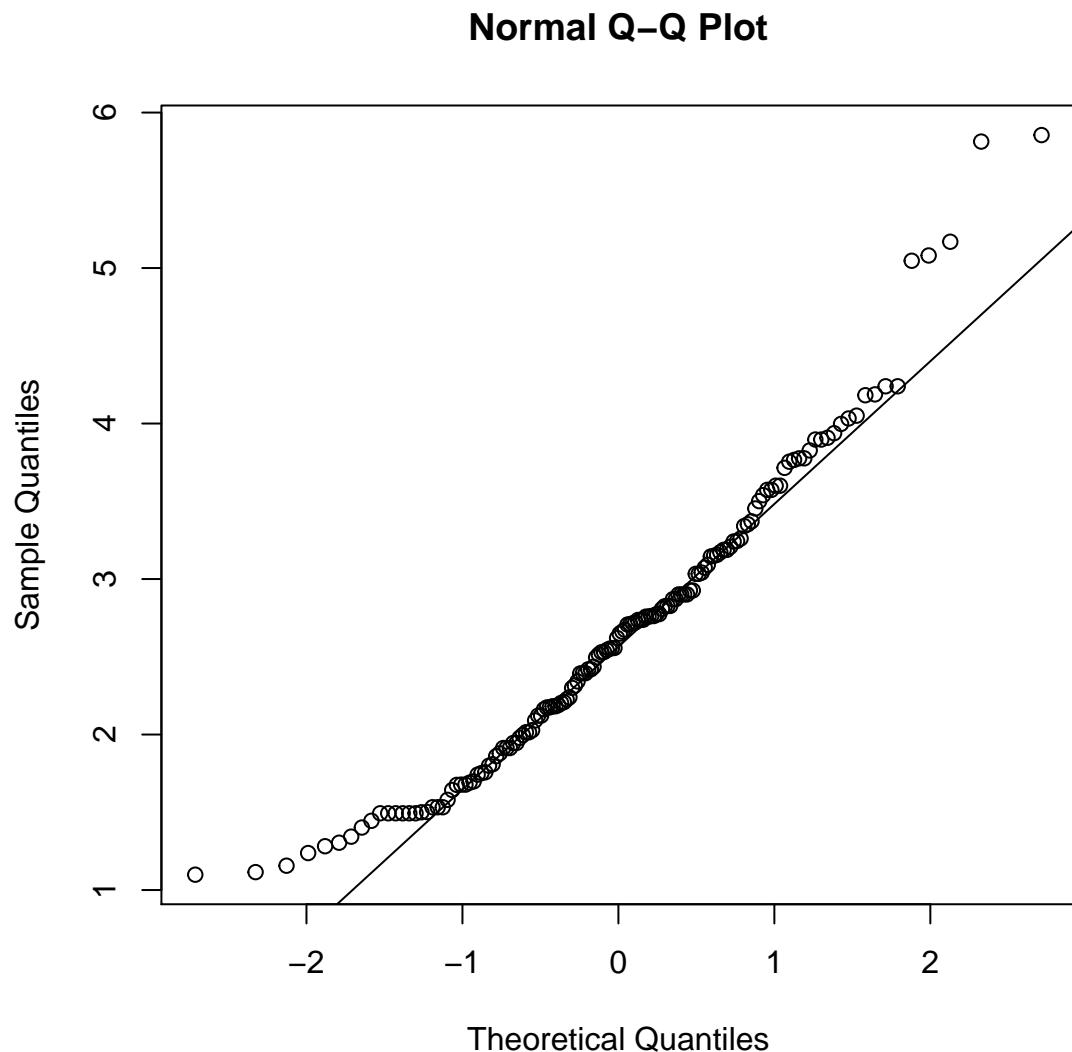
This is done to check the multivariate normality of the data. If the data are normally distributed, we may use maximum likelihood (ML) estimation method for the CFA. In `lavaan`, we have a number of

alternative estimation methods (the full list is available at <http://lavaan.ugent.be/tutorial/est.html> or by typing `?lavOptions`). Two common alternatives are:

1. **MLR** (robust ML), suitable for complete and incomplete, non-normal data (Rosseel, 2018).
2. **WLSMV** (robust weighted least squares), suitable for categorical response options (e.g. dichotomous, polynomous, ordinal (Brown, 2015))

```
mardia(data.cfa)
```

```
## Call: mardia(x = data.cfa)
##
## Mardia tests of multivariate skew and kurtosis
## Use describe(x) the to get univariate tests
## n.obs = 150    num.vars =  8
## b1p =  11.37   skew =  284.27 with probability =  2.3e-15
## small sample skew =  291.24 with probability =  3.3e-16
## b2p =  96.9    kurtosis =  8.18 with probability =  2.2e-16
```



the data are not multivariate normal (`kurtosis > 5, P < 0.05`). We will use **MLR** in our analysis.

3.2 Step 1

Specify the measurement model

Specify the measurement model according to `lavaan` syntax.

```
model = "
PA1 =~ Q4 + Q5 + Q6 + Q7 + Q11
PA2 =~ Q8 + Q9 + Q10
"
```

`=~` indicates “measured by”, thus the items represent the factor.

By default, `lavaan` will correlate PA1 and PA2 (i.e. `PA1 ~~ PA2`), somewhat similar to oblique rotation in EFA. `~~` means “correlation”. We will use `~~` when we add correlated errors later.

3.3 Step 2

Fit the model

Here, we fit the specified model. By default, marker indicator variable approach¹ is used in `lavaan` to scale a factor². We use `MLR` as the estimation method.

```
cfa.model = cfa(model, data = data.cfa, estimator = "MLR")
# cfa.model = cfa(model, data = data.cfa, std.lv = 1) # factor variance = 1
summary(cfa.model, fit.measures = T, standardized = T)

## lavaan 0.6-3 ended normally after 21 iterations
##
## Optimization method                           NLMINB
## Number of free parameters                   17
##
## Number of observations                      150
##
## Estimator                                    ML      Robust
## Model Fit Test Statistic                  37.063   27.373
## Degrees of freedom                         19       19
## P-value (Chi-square)                      0.008   0.096
## Scaling correction factor                 1.354
##           for the Yuan-Bentler correction (Mplus variant)
##
## Model test baseline model:
##
## Minimum Function Test Statistic          453.795   325.195
## Degrees of freedom                         28       28
## P-value                                  0.000   0.000
##
## User model versus baseline model:
##
## Comparative Fit Index (CFI)              0.958   0.972
## Tucker-Lewis Index (TLI)                 0.937   0.958
##
```

¹The regression weight of an item from a factor is fixed to 1. Another approach in CFA is to fix the factor variance to 1 (Brown, 2015).

²The latent variable (factor) is an unobserved variable, thus it has to be scaled by a method to define its metric/unit of measurement. This is done by fixing either the item regression weight or the factor variance to 1.

```

## Robust Comparative Fit Index (CFI) 0.973
## Robust Tucker-Lewis Index (TLI) 0.960
##
## Loglikelihood and Information Criteria:
##
## Loglikelihood user model (H0) -1566.019 -1566.019
## Scaling correction factor 1.140
## for the MLR correction
## Loglikelihood unrestricted model (H1) -1547.487 -1547.487
## Scaling correction factor 1.253
## for the MLR correction
##
## Number of free parameters 17 17
## Akaike (AIC) 3166.037 3166.037
## Bayesian (BIC) 3217.218 3217.218
## Sample-size adjusted Bayesian (BIC) 3163.416 3163.416
##
## Root Mean Square Error of Approximation:
##
## RMSEA 0.080 0.054
## 90 Percent Confidence Interval 0.040 0.118 0.000 0.091
## P-value RMSEA <= 0.05 0.098 0.397
##
## Robust RMSEA 0.063
## 90 Percent Confidence Interval 0.000 0.112
##
## Standardized Root Mean Square Residual:
##
## SRMR 0.079 0.079
##
## Parameter Estimates:
##
## Information Observed
## Observed information based on Hessian
## Standard Errors Robust.huber.white
##
## Latent Variables:
## Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## PA1 =~
## Q4 1.000 0.952 0.814
## Q5 0.660 0.092 7.218 0.000 0.629 0.624
## Q6 0.810 0.090 9.010 0.000 0.771 0.708
## Q7 0.916 0.086 10.641 0.000 0.872 0.735
## Q11 0.533 0.093 5.719 0.000 0.507 0.544
## PA2 =~
## Q8 1.000 0.653 0.655
## Q9 1.347 0.156 8.654 0.000 0.880 0.844
## Q10 1.436 0.199 7.206 0.000 0.938 0.856
##
## Covariances:
## Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## PA1 ~~
## PA2 0.077 0.075 1.035 0.301 0.124 0.124
##

```

```

## Variances:
##           Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## .Q4        0.460   0.089   5.182   0.000   0.460   0.337
## .Q5        0.620   0.086   7.178   0.000   0.620   0.611
## .Q6        0.590   0.101   5.836   0.000   0.590   0.498
## .Q7        0.647   0.125   5.181   0.000   0.647   0.460
## .Q11       0.611   0.077   7.934   0.000   0.611   0.704
## .Q8        0.567   0.101   5.628   0.000   0.567   0.570
## .Q9        0.312   0.094   3.325   0.001   0.312   0.287
## .Q10       0.321   0.106   3.046   0.002   0.321   0.267
## PA1        0.906   0.137   6.587   0.000   1.000   1.000
## PA2        0.427   0.106   4.009   0.000   1.000   1.000

```

Results

Read the results marked as **Robust**. These represent the results of **MLR**.

To interpret the results, we must looks at

1. Overall model fit - by fit indices.
 2. Localized areas of misfit
 - Residuals.
 - Modification indices.
 3. Parameter estimates
 - Factor loadings (**Std.all** column under **Latent Variables** table).
 - Factor correlations (**Std.all** column under **Covariances** table).
1. Fit indices.

The following are a number of selected fit indices and the recommended cut-off values (Brown, 2015; Schreiber, Nora, Stage, Barlow, & King, 2006),

Category	Fit index	Cut-off
Absolute fit	χ^2	$P > 0.05$
	Standardized root mean square (SRMR)	≤ 0.08
Parsimony correction	Root mean square error of approximation (RMSEA)	and its 90% CI ≤ 0.08 , CFit $P > 0.05$
Comparative fit	Comparative fit index (CFI)	≥ 0.95
	Tucker-Lewis index (TLI)	

2. Localized areas of misfit (Brown, 2015)

- Residuals

Residuals are the difference between the values in the sample and model-implied variance-covariance matrices.

Standardized residuals (SRs) $> |2.58|$ indicate the standardized discrepancy between the matrices.

- Modification indices (MIs)

A modification index indicates the expected parameter change if we include a particular specification in the model (i.e. a constrained/fixed parameter is freely estimated, e.g. by correlating between errors of Q1 and Q2).

Specifications with MIs > 3.84 should be investigated.

3. Parameter estimates

- Factor loadings (FLs) (**Std.all** column under **Latent Variables** table).

The guideline for EFA is applicable also to CFA. For example, $FLs \geq 0.5$ are practically significant. In addition, the P -values of the FLs must be significant (at $\alpha = 0.05$).

Also look for out-of-range values. FLs should be in range of 0 to 1 (absolute values), thus values > 1 are called *Heywood cases* or *offending estimates* (Brown, 2015)

- Factor correlations (Std.all column under Covariances table).

Similar to EFA, a factor correlation must be < 0.85 , which indicates that the factors are distinct. A correlation > 0.85 indicates multicollinearity problem. Also look for out-of-range values. Factor correlations should be in range of 0 to 1 (absolute values).

In addition, when a model has Heywood cases, the solution is not acceptable. The variance-covariance matrix (of our data) could be *non-positive definite* i.e. the matrix is not invertible for the analysis.

In our output:

Fit indices,

```
## Number of observations                                150
##
## Estimator                                         ML   Robust
## Model Fit Test Statistic                         37.063 27.373
## Degrees of freedom                               19     19
## P-value (Chi-square)                            0.008  0.096
## Scaling correction factor                      1.354
##           for the Yuan-Bentler correction (Mplus variant)
##
## Comparative Fit Index (CFI)                     0.958  0.972
## Tucker-Lewis Index (TLI)                         0.937  0.958
##
## RMSEA                                            0.080  0.054
## 90 Percent Confidence Interval                 0.040  0.118  0.000  0.091
## P-value RMSEA <= 0.05                           0.098  0.397
##
## Robust RMSEA                                      0.063
## 90 Percent Confidence Interval                  0.000  0.112
##
## Standardized Root Mean Square Residual:
##
## SRMR                                           0.079  0.079
```

The model has good model fit based on all indices, with the exception of the upper 90% CI of robust RMSEA = 0.112. Please note there is no CFit P -value for robust RMSEA.

FLs and factor correlation,

```
## Latent Variables:
##                               Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## PA1 =~
##   Q4          1.000
##   Q5          0.660  0.092  7.218  0.000  0.629  0.624
##   Q6          0.810  0.090  9.010  0.000  0.771  0.708
##   Q7          0.916  0.086 10.641  0.000  0.872  0.735
##   Q11         0.533  0.093  5.719  0.000  0.507  0.544
## PA2 =~
##   Q8          1.000
##   Q9          1.347  0.156  8.654  0.000  0.880  0.844
##   Q10         1.436  0.199  7.206  0.000  0.938  0.856
```

```

## Covariances:
##                               Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## PA1 ~~
## PA2             0.077    0.075   1.035   0.301   0.124   0.124

```

Remember to read the results down the `Std.all` column. All FLs > 0.5 and the factor correlation < 0.85. There is no problem with the item quality and the factors are distinct.

In addition, to obtain the standardized results with SE,

```
standardizedSolution(cfa.model) # standardized, to view the SE of FL
```

```

##   lhs op rhs est.std   se      z pvalue ci.lower ci.upper
## 1 PA1 =~ Q4  0.814 0.041 20.017 0.000   0.735  0.894
## 2 PA1 =~ Q5  0.624 0.068  9.117 0.000   0.490  0.758
## 3 PA1 =~ Q6  0.708 0.060 11.867 0.000   0.591  0.825
## 4 PA1 =~ Q7  0.735 0.059 12.523 0.000   0.620  0.850
## 5 PA1 =~ Q11 0.544 0.080  6.781 0.000   0.387  0.702
## 6 PA2 =~ Q8  0.655 0.067  9.712 0.000   0.523  0.788
## 7 PA2 =~ Q9  0.844 0.048 17.646 0.000   0.751  0.938
## 8 PA2 =~ Q10 0.856 0.048 17.670 0.000   0.761  0.951
## 9  Q4 ~~ Q4  0.337 0.066  5.078 0.000   0.207  0.467
## 10 Q5 ~~ Q5  0.611 0.085  7.156 0.000   0.444  0.778
## 11 Q6 ~~ Q6  0.498 0.085  5.894 0.000   0.333  0.664
## 12 Q7 ~~ Q7  0.460 0.086  5.324 0.000   0.290  0.629
## 13 Q11 ~~ Q11 0.704 0.087  8.049 0.000   0.532  0.875
## 14 Q8 ~~ Q8  0.570 0.088  6.446 0.000   0.397  0.744
## 15 Q9 ~~ Q9  0.287 0.081  3.550 0.000   0.128  0.445
## 16 Q10 ~~ Q10 0.267 0.083  3.226 0.001   0.105  0.430
## 17 PA1 ~~ PA1 1.000 0.000     NA     NA   1.000  1.000
## 18 PA2 ~~ PA2 1.000 0.000     NA     NA   1.000  1.000
## 19 PA1 ~~ PA2 0.124 0.113  1.099 0.272  -0.098  0.346

```

and, to obtain the unstandardized results with 95% CI,

```
parameterEstimates(cfa.model) # unstandardized, to view the 95% CI
```

```

##   lhs op rhs   est    se      z pvalue ci.lower ci.upper
## 1 PA1 =~ Q4 1.000 0.000     NA     NA   1.000  1.000
## 2 PA1 =~ Q5 0.660 0.092  7.218 0.000   0.481  0.840
## 3 PA1 =~ Q6 0.810 0.090  9.010 0.000   0.634  0.986
## 4 PA1 =~ Q7 0.916 0.086 10.641 0.000   0.748  1.085
## 5 PA1 =~ Q11 0.533 0.093  5.719 0.000   0.350  0.716
## 6 PA2 =~ Q8 1.000 0.000     NA     NA   1.000  1.000
## 7 PA2 =~ Q9 1.347 0.156  8.654 0.000   1.042  1.652
## 8 PA2 =~ Q10 1.436 0.199  7.206 0.000   1.046  1.827
## 9  Q4 ~~ Q4 0.460 0.089  5.182 0.000   0.286  0.633
## 10 Q5 ~~ Q5 0.620 0.086  7.178 0.000   0.451  0.789
## 11 Q6 ~~ Q6 0.590 0.101  5.836 0.000   0.392  0.788
## 12 Q7 ~~ Q7 0.647 0.125  5.181 0.000   0.402  0.891
## 13 Q11 ~~ Q11 0.611 0.077  7.934 0.000   0.460  0.762
## 14 Q8 ~~ Q8 0.567 0.101  5.628 0.000   0.369  0.764
## 15 Q9 ~~ Q9 0.312 0.094  3.325 0.001   0.128  0.495
## 16 Q10 ~~ Q10 0.321 0.106  3.046 0.002   0.115  0.528
## 17 PA1 ~~ PA1 0.906 0.137  6.587 0.000   0.636  1.175
## 18 PA2 ~~ PA2 0.427 0.106  4.009 0.000   0.218  0.635
## 19 PA1 ~~ PA2 0.077 0.075  1.035 0.301  -0.069  0.224

```

Localized areas of misfit,

```
mi = modificationIndices(cfa.model)
subset(mi, mi > 3.84)

##   lhs op rhs      mi    epc sepc.lv sepc.all sepc.nox
## 20 PA1 =~ Q8 10.264 0.244    0.232    0.233    0.233
## 24 PA2 =~ Q5  8.359 0.332    0.217    0.215    0.215
## 37 Q5 ~~ Q11 6.301 0.144    0.144    0.234    0.234
## 55 Q9 ~~ Q10 10.264 2.325    2.325    7.346    7.346

sr = residuals(cfa.model, type="standardized"); sr

## $type
## [1] "standardized"
##
## $cov
##      Q4      Q5      Q6      Q7      Q11     Q8      Q9      Q10
## Q4  0.000
## Q5  0.007  0.000
## Q6 -0.159 -0.361  0.000
## Q7  0.214 -1.113  0.594  0.000
## Q11 -0.075  1.568 -0.878 -0.657  0.000
## Q8   1.599  3.397  1.464  1.268  1.139  0.000
## Q9  -0.940  2.065  0.131 -1.397  0.936 -0.116  0.000
## Q10 -1.092  1.495 -1.121 -1.319  0.354 -0.042  0.065  0.000
```

There are four suggested specifications with MIs > 3.84 . We may ignore $PA1 \sim Q8$ and $PA2 \sim Q5$ based on content, because it is not justifiable to allow these two items specified under other factors. $Q9 \sim Q10$ is justifiable, based on the wording “is important”. But $Q5 \sim Q11$ is not justifiable.

$Q5$ has an SR > 2.58 with $Q8$ (SR = 3.397). So we may focus on $Q5$. Avoid $Q8$ because there are only three items in the factor and FL $Q8 > Q5$.

3.4 Step 3

Whenever the model do not fit well, we must revise the model. To do so, we must look for the causes of the poor fit to the data. The causes in CFA could be:

1. Item – the item has low FL (< 0.3), is specified to load on wrong factor or has cross-loading issue.
2. Factor – the factors have multicollinearity problem (correlation > 0.85), or the presence of redundant factors in a model. This can detected by residuals and MIs.
3. Correlated error (method effect) – some items are similarly worded (e.g. “I like ...”, “I believe...”) or have almost similar meaning/content. This is usually detected by residuals and MIs.
4. Improper solution – the solution with Heywood cases. It could be because the specified model is not supported by the data and the misspecification could be a combination of all the first three causes listed above. A small sample may also lead to improper solution.

The problems might not surface if a proper EFA is done in the first place and the model is theoretically sound.

Model-to-model comparison following revision is done based on:

1. χ^2 difference
 - for nested³ models only.

³model with same number of items, but with different model specifications e.g. number of factors

2. AIC (Akaike Information Criterion) and BIC (Bayesian Information Criterion)

- for nested and unnested models.
- an improvement in the model is shown as a reduction in AIC and BIC values (Brown, 2015). Better model = Smaller AIC/BIC.

Model revision

Revision 1: Based on MI, Q9 ~ Q10?

Both from PA2, reasonable by the wording of the questions.

```
model1 =
PA1 =~ Q4 + Q5 + Q6 + Q7 + Q11
PA2 =~ Q8 + Q9 + Q10
Q9 ~~ Q10
"
cfa.model1 = cfa(model1, data = data.cfa, estimator = "MLR")

## Warning in lav_object_post_check(object): lavaan WARNING: some estimated ov variances are
## negative

summary(cfa.model1, fit.measures=T, standardized=T)
```

Take note of the warning message above! It points out to problem(s) in our model. Here we have *negative variance*. Remember that variance is the square of standard deviation, thus it is impossible to have a negative variance!

```
## Number of observations                                150
##
## Estimator                                         ML    Robust
## Model Fit Test Statistic                         26.487 19.771
## Degrees of freedom                               18     18
## P-value (Chi-square)                            0.089  0.346
## Scaling correction factor                      1.340
##      for the Yuan-Bentler correction (Mplus variant)
##
## Comparative Fit Index (CFI)                     0.980  0.994
## Tucker-Lewis Index (TLI)                         0.969  0.991
##
## RMSEA                                           0.056  0.026
## 90 Percent Confidence Interval          0.000  0.099  0.000  0.073
## P-value RMSEA <= 0.05                           0.376  0.754
##
## Robust RMSEA                                    0.030
## 90 Percent Confidence Interval          0.000  0.091
##
## Standardized Root Mean Square Residual:
##
## SRMR                                         0.058  0.058
```

The upper 90% CI of RMSEA is smaller, but

```
## Latent Variables:
##                               Estimate Std.Err  z-value P(>|z|) Std.lv Std.all
## PA1 =~
##   Q4                  1.000
##   Q5                 0.663   0.089   7.470   0.000   0.950   0.813
##   Q6                 0.811   0.091   8.876   0.000   0.771   0.708
```

```

##    Q7          0.923   0.089   10.409   0.000   0.877   0.739
##    Q11         0.528   0.092    5.729   0.000   0.502   0.538
##  PA2 =~
##    Q8          1.000
##    Q9          0.247   0.314    0.786   0.432   0.374   0.359
##    Q10         0.267   0.332    0.803   0.422   0.404   0.369
## Covariances:
##              Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## .Q9 ~~
## .Q10          0.678   0.198    3.419   0.001   0.678   0.683
## PA1 ~~
## PA2          0.267   0.096    2.787   0.005   0.185   0.185

```

we have a serious Heywood case here! Q8 FL = 1.522. Thus this solution is not acceptable.

Revision 2: Remove Q5?

```

model2 =
PA1 =~ Q4 + Q6 + Q7 + Q11
PA2 =~ Q8 + Q9 + Q10
"
cfa.model2 = cfa(model2, data = data.cfa, estimator = "MLR")
summary(cfa.model2, fit.measures=T, standardized=T)

```

```

## Number of observations                               150
##
## Estimator                                         ML      Robust
## Model Fit Test Statistic                         20.451   14.467
## Degrees of freedom                                13       13
## P-value (Chi-square)                            0.085   0.342
## Scaling correction factor                      1.414
## for the Yuan-Bentler correction (Mplus variant)
##
## Comparative Fit Index (CFI)                     0.979   0.994
## Tucker-Lewis Index (TLI)                        0.966   0.990
##
## RMSEA                                           0.062   0.027
## 90 Percent Confidence Interval                 0.000   0.111   0.000   0.079
## P-value RMSEA <= 0.05                          0.314   0.704
##
## Robust RMSEA                                     0.033
## 90 Percent Confidence Interval                  0.000   0.104
##
## Standardized Root Mean Square Residual:
##
## SRMR                                         0.063   0.063

```

The upper 90% CI of RMSEA has reduced from 0.112 to 0.104.

```

## Latent Variables:
##              Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##  PA1 =~
##    Q4          1.000
##    Q6          0.830   0.099    8.391   0.000   0.781   0.718
##    Q7          0.960   0.100    9.597   0.000   0.904   0.762
##    Q11         0.504   0.091    5.547   0.000   0.474   0.509
##  PA2 =~

```

```

##      Q8          1.000
##      Q9          1.351  0.155  8.692  0.000  0.651  0.653
##      Q10         1.444  0.202  7.143  0.000  0.880  0.844
## Covariances:
##                  Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## PA1 ~~
## PA2          0.048   0.068  0.705  0.481  0.078  0.078

The FLs and factor correlation are acceptable. No Heywood's case.

mi2 = modificationIndices(cfa.model2)
subset(mi2, mi > 3.84)

##   lhs op rhs    mi   epc sepc.lv sepc.all sepc.nox
## 18 PA1 =~ Q8 9.707 0.241   0.227   0.228   0.228
## 45 Q9 ~~ Q10 9.707 3.661   3.661  11.615  11.615

sr2 = residuals(cfa.model2, type="standardized"); sr2

## $type
## [1] "standardized"
##
## $cov
##      Q4      Q6      Q7      Q11     Q8      Q9      Q10
## Q4  0.000
## Q6 -0.231  0.000
## Q7 -0.089  0.390  0.000
## Q11 0.594 -0.408 -0.461  0.000
## Q8  1.970  1.759  1.579  1.407  0.000
## Q9 -0.495  0.488 -1.183  1.325 -0.115  0.000
## Q10 -0.524 -0.771 -1.043  0.650 -0.051  0.077  0.000

```

There is no SR > 2.58.

So we may stop at **model2**, although the upper 90% CI of RMSEA is still > 0.08, but there is no more localized areas of misfit by SR.

Model-to-model comparison

Because **model2** is not nested in **model**, we compare mainly by AIC and BIC, and additionally by χ^2 difference (in our case scaled χ^2 difference),

```
anova(cfa.model, cfa.model2, method = "satorra.bentler.2010")
```

```

## Scaled Chi Square Difference Test (method = "satorra.bentler.2001")
##
##           Df  AIC   BIC  Chisq Chisq diff Df diff Pr(>Chisq)
## cfa.model2 13 2791 2836.2 20.451
## cfa.model  19 3166 3217.2 37.063      13.562       6   0.03493 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Clearly, the AIC and BIC are reduced (**model2** [without Q5] vs **model** [with Q5]). The χ^2 difference is significant, which indicates an improvement in model fit.

4 Construct reliability

Raykov's rho

Raykov's rho is one of the reliability indices applicable to CFA. It takes into account the correlated errors. Construct reliability ≥ 0.7 (Hair, Black, Babin, & Anderson, 2010) is acceptable.

Look at the `omega` row in the output,

```
rel.model2 = reliability(cfa.model2)
print(rel.model2, digits = 3)
```

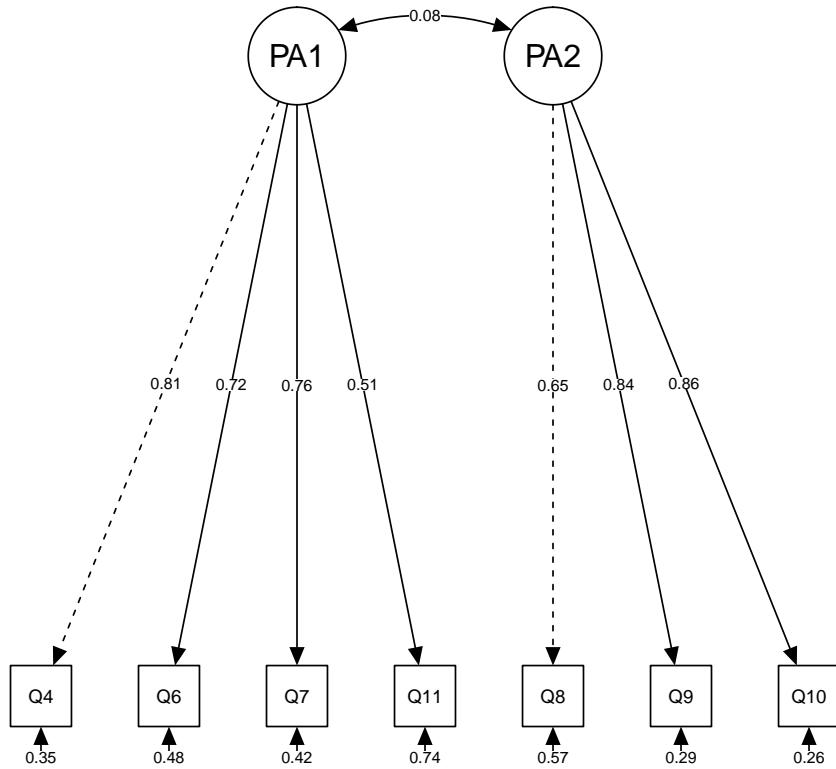
```
##          PA1    PA2 total
## alpha   0.792  0.826  0.723
## omega   0.808  0.836  0.829
## omega2  0.808  0.836  0.829
## omega3  0.809  0.836  0.793
## avevar  0.526  0.634  0.570
```

Raykov's rho (the `omega`): **PA1** = 0.808, **PA2** = 0.836. Both factors are reliable.

5 Path diagram

A CFA model can be nicely presented in the form of path diagram.

```
semPaths(cfa.model2, 'path', 'std', style = 'lisrel',
          edge.color = 'black', intercepts = F)
```



6 Results presentation

In the report, you must include a number of important statements and results pertaining to the CFA,

1. The estimation method e.g. ML, MLR, WLSMV etc.
2. The model specification and the theoretical background supporting the model.
3. Details about the selected fit indices, residuals, MIIs, FLs and factor correlations and the accepted cut-off values.
4. Detailed comments on the fit and parameters of the tested models. This is usually done in reference to summary tables.
5. Details about the revision process, i.e. item deletion, addition of correlated errors or any other modifications and the effects on the model fit. Also mention the reasons e.g. high SRs, low FLs etc.
6. Summary tables, which outlines the model fit indices, model comparison, FLs, communalities, Raykov's rho, and factor correlations.
7. The path diagram (most of the time, of the final model). This may be requested by some journals.

Fit indices of the models.

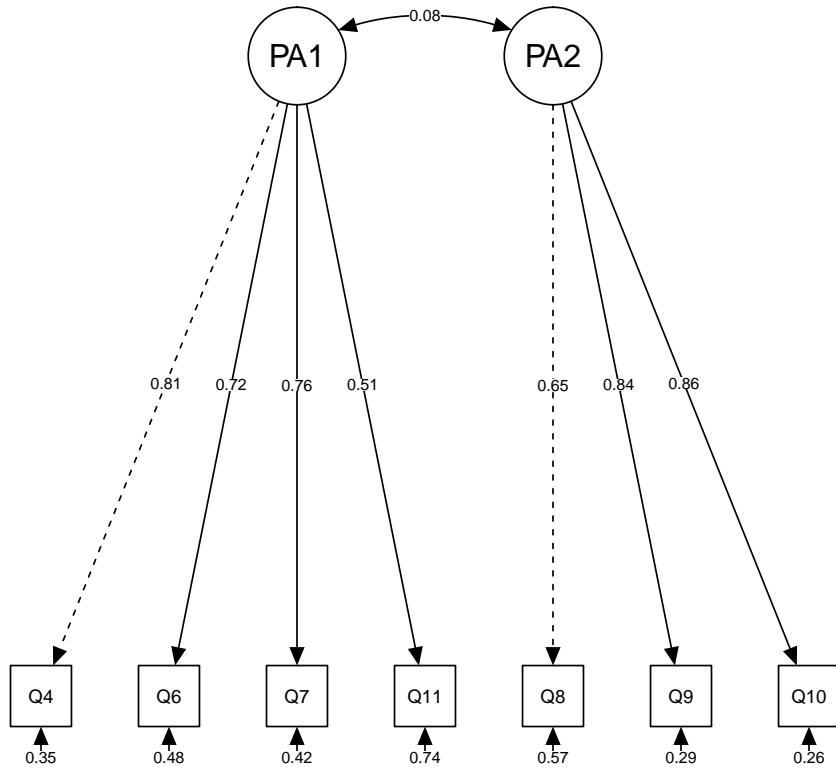
Model	$\chi^2(\text{df})$	P	$\chi^2_{diff}(\text{df})$	P	SRMR	RMSEA	90%		AIC	BIC	
							CI	CFI			
Model 2	27.4(19)	0.096	-		0.079	0.063	0.000, 0.112	0.973	0.960	3166	3217
Model 2	14.5(13)	0.342	13.6 (6)	0.035	0.063	0.033	0.000, 0.104	0.994	0.990	2791	2836

Factor loadings and reliability of Model 2.

Factor	Item	Factor loading	Raykov's rho
Affinity	Q4	0.806	0.808
	Q6	0.718	
	Q7	0.762	
	Q11	0.509	
Importance	Q8	0.653	0.836
	Q9	0.844	
	Q10	0.858	

Factor correlation: Affinity \leftrightarrow Importance $r = 0.078$.

The path diagram of Model 2.



References

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